Does land conservation raise property taxes? Evidence from New England cities and towns

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A B S T R A C T

Protected lands provide high ecological and social value, yet a perception that land protection erodes local property tax bases and shifts tax burdens to other owners creates barriers for new protection. We investigate the impact of land protection on local property tax rates using panel data from more than 1400 towns and cities in New England between 1990 and 2015, including both ownership and easement-based protection. We find that on average, new protection results in a small increase in annual tax rates, with an expected change of $0.0231 per $1000 of value for a 1 percentage point increase in the percent of municipal area protected. This corresponds to an increase in a homeowner’s annual property tax bill of $1.16 per $100,000 of value for 100 acres of new land protection. We do not find that taxes continue to increase over time or reduce municipal expenditures. However, for towns that are growing slowly, have lower household incomes, or use municipal land protection, we estimate greater tax bill increases of up to $10 for each $100,000 of value. These results provide evidence that land protection does not have a substantial impact on property taxes, but also highlight the importance of maintaining and expanding public compensation mechanisms, such as payments in lieu of taxes, where expected burdens from new protection may be greater.

1. Introduction

Protected land provides multiple ecological and social benefits including carbon sequestration, habitat for a diverse set of plant and animal species, watersheds functioning, preservation of prime agricultural soils, and space for recreation and cultural preservation (Dinerstein et al., 2019; Watson et al., 2014; Brauman et al., 2007). In the U.S., land protection has expanded rapidly in recent decades, motivated by the continued loss of open space to development (Nelson et al., 2007; Kotchen and Powers 2006). This increase has been facilitated by state and federal funding (The Trust for Public Lands 2021; Stubbs 2020), shifts in land ownership (Meyer et al., 2014), tax incentives (Parker and Thurman 2018), and funding for open space protection through local referenda (Lang 2018). Substantial federal funding is currently available for new land protection through the reauthorized Land and Water Conservation Fund (National Park Service 2020) and the Farm Bill (Stubbs 2020) as well as the recently passed Inflation Reduction Act (Congressional Research Service 2022) and infrastructure bills (Rigley 2021).

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The surge in funding has renewed debate about the fiscal implications of land protection. Land that is protected through conservation restrictions or under ownership by public and non-profit organizations is frequently tax-exempt or taxed at lower rates than developed or unprotected land. Critics of new land protection worry that it will erode local property tax bases and result in higher property tax rates for other landowners (Brandon 2021; Ricketts 2021; Rule 2019; Neuman 2018; LePage 2018; McWhirter 2014). These concerns have fueled opposition to both local land protection efforts (LePage 2018) and national conservation initiatives including President Biden’s 30-by-30 conservation target (Brandon 2021). Proponents of land protection counter by arguing that open space pays for itself because it requires less in municipal expenditures than it contributes to revenues, or because it creates recreation-based economic opportunities and amenity values that increase the property tax base (e.g., Davis et al., 2018; The Trust for Public Land 2007).

Despite the crucial role that local property taxes play in funding public goods, few studies have quantified the possible causal impacts of land protection on property tax rates to inform this debate (Vandegrift and Lahr 2011; King and Anderson 2004). Indeed, estimating these relationships is difficult due to the potentially endogenous nature of land protection. For example, communities that engage in more land protection may tend to have greater wealth, past histories of land protection, or more to gain from recreation-based economies. Alternatively, communities with more land protection may be more rural and have fewer development opportunities.

To overcome these potential selection bias concerns, we exploit plausibly exogenous changes in land protection over time within municipalities. Specifically, we estimate the effects of new land protection on changes in local property tax rates using municipal-level panel data from 1990 to 2015 across five New England states. Our data combines information on property tax rates, property tax levies, and taxable property value for more than 1400 municipalities in the region (also referred to as New England’s “towns and cities”). We match these with detailed spatial data on new land protection over time assembled by the Harvard Forest and the Highstead Foundation, which follow the Protected Areas Database of the United States (PAD-US) in defining protected lands as areas “dedicated to the preservation of biological diversity and to other natural, recreational and cultural uses, managed for these purposes through legal or other effective means” (U.S. Geological Survey, 2022).

Our regression model estimates changes in property tax rates as a function of lagged changes in land protection, with controls for state-by-time-period fixed effects, core-based statistical area trends, and lagged changes in the labor force, unemployment, and property tax base growth. This approach differences out unobserved time-invariant characteristics of municipalities and reduces serial correlation. The key identifying assumption inherent to this empirical strategy is that the timing of new protection within municipalities, conditional on these controls, is exogenous to potential outcomes. This is plausible because protection in the region is the result of decentralized and uncoordinated actions by more than 350 separate land trusts, hundreds of local governments, and multiple state and federal agencies. Land trusts overlap in their spatial jurisdictions and missions (Wildlands and Woodlands 2021; Foster et al., 2017; Labich 2015), and many local protections have been driven by the efforts of just a few committed individuals. In addition, opportunities for protection often occur when there is a generational shift within families due to retirements, health shocks, or deaths (Markowski-Lindsay et al., 2017; Bigelow et al., 2016). The nature of these protection processes within New England creates considerable randomness in the extent and timing of new land protection.

We estimate average effects and test for differential impacts across land protection types and by several key characteristics of the communities where land protection is occurring. We also examine fiscal outcomes (municipal revenues, expenditures) for the states with available data in order to study broader fiscal responses to land protection and we estimate a model with lagged protection to investigate longer-term impacts on tax rates.

We find that on average, new land protection has had small impacts on property tax rates. We estimate an expected increase in annual property tax rates of $0.0231 per $1000 of value that is attributable to a one percentage point increase in new municipal area protected annually. For an annual increase in new land protection of 100 acres, this property tax rate change translates to an annual property tax bill increase of just $1.16 per $100,000 of value, or $3.00 for an owner of a typical New England home. We do not find evidence that municipalities collect less revenue or reduce expenditures on public goods as a result of land protection, or that there are additional longer-term impacts on the tax rate.\(^1\)

While average impacts are small, we do observe important heterogeneity in property tax bill impacts by land protection type and local characteristics, with annual increases of up to $10 per $100,000 of home valuation. The types of towns that are associated with greater estimated impacts are those that are growing slowly, have lower median incomes, have less land enrolled in current use programs that allow reduced taxes for agriculture and forestry, or have fewer second homes. We also find suggestive evidence for larger tax increases associated with land protection conducted by municipalities, particularly in towns and cities with smaller property tax bases or slower growth.

Overall, our findings indicate that the property tax rate changes due to land protection are generally not substantial, particularly in comparison to the magnitude of changes that residents may experience for capital projects such as new buildings or increases in municipal staff. Yet the heterogeneity in impacts highlights the reality that some communities may be at higher risk for greater tax burdens. These differences emphasize a need for public compensation mechanisms, such as state and federal payments in lieu of taxes, that can assist communities engaging in land protection, and provide a rationale for targeting these programs to the types of communities that may be most impacted by new land protection.

\(^1\) Valued at $259,045 in 2015 dollars, according to Zillow Home Value Index for 1996–2015 that reflects the mean value of single-family homes in the 35th to 65th percentile range of home values (Zillow Research 2019).
2. Conceptual framework and contributions to the literature

2.1. Conceptual framework

Economic theory suggests that the impacts of land protection on local property tax rates could be either positive or negative in net (Wu et al., 2016; Wu 2014; King and Anderson 2004; Geoghegan et al., 2003). This can be illustrated in a simple way by considering the relationship between the property tax rate, municipal property tax base, expenditures and revenues (assuming a balanced budget and a single tax rate):

\[ \frac{\text{Property Tax Rate}}{\text{Property Tax Base}} = \frac{\text{Property Tax Levy}}{\text{Property Tax Base}} - \frac{\text{Other Revenue}}{\text{Property Tax Base}} = \frac{\text{Total Expenditures}}{\text{Property Tax Base}} \]  

(1)

The property tax rate is determined by the funds needed to cover expected expenditures, less other available sources of revenue, divided by the value of the taxable property in a municipality. This amount of property tax revenue is raised via the property tax levy (revenue raised from the property tax).

A common concern is that newly protected land will have lower taxable value or will be removed from the tax rolls altogether if the owner is tax exempt, which would reduce the total municipal property tax base. If expected expenditures and other revenues are constant, such a reduction in the property tax base would require an increase in the property tax rate to raise the same total property tax revenues.

On the other hand, land protection can create amenity value which may boost nearby property values or the desirability of a municipality as a place to live and property values in the municipality as a whole (Lang 2018; Vandegrift and Lahr 2011; Anderson and West 2006; Irwin 2002). Amenity-driven property value growth can increase the local property tax base and reduce the need to raise property tax rates or even lead to lower tax rates. Land protection may also attract new amenity-based development in nearby areas if developable land is available, potentially growing the tax base (Davis and Hansen 2011; Wade and Theobald 2010; Radeloff et al., 2010). In addition, land protection can create opportunities for recreation-based economic activity (e.g., Walls et al., 2020; Sims et al., 2019; Chen et al., 2016; Rasker et al., 2013), increasing other revenue sources for a municipality. Finally, the cost of services may be lower for undeveloped than developed land. Preventing land from being developed may reduce current or future revenues but may also keep expected expenditures low by limiting service needs, reducing pressures on property tax rate growth over time (Murray and Catanzaro 2019; Kotchen and Schulte 2009).

Our study seeks to understand the overall impact of land protection on property tax rates, which is an empirical question that depends on the magnitudes of these individual channels through which protection can affect property tax rates.

2.2. Contributions to the literature

Prior research corroborates each of these possible channels, although only a limited set of previous studies directly explores the impacts of land protection on property tax rates. Much of the related prior literature in economics establishes the importance of amenity values. Research on how the value of open space is capitalized into property values includes studies at the parcel (e.g., Chamblee et al., 2011; Anderson and West 2006; Geoghegan et al., 2003; Irwin 2002; Thorndes 2002), zip code (Lang 2018), and municipal (Vandegrift and Lahr 2011) levels.

Hedonic studies of open space impacts on property values consistently find positive localized impacts. Studies of residential property values report impacts ranging from 0.05% to 1.87% (Anderson and West 2006; Geoghegan et al., 2003; Irwin 2002), while studies of vacant land sales show that land values can increase by as much as 19–46% due to adjacency to protected land (Thorndes 2002; Chamblee et al., 2011). At the zip code and municipal levels, Lang (2018) and Vandegrift and Lahr (2011) also observe positive impacts of open space expenditures on home values.

Previous literature also finds important heterogeneity in amenity values. Studies have demonstrated a positive relationship between willingness to pay for open space and income (Earnhart 2006; Breffle et al., 1998), and have found increasing demand for open space as a function of income in the context of open space referenda (Nelson et al., 2007; Kline 2006; Kotchen and Powers 2006), environmental ballot measures (Kahn and Matusasa 1997), and municipal open space acquisitions (Bates and Santerre 2001). There is also evidence that the amenity value of nearby open space increases with neighborhood income (Anderson and West 2006). Accordingly, land protection may disproportionately boost property values in high income municipalities. Additionally, higher income municipalities may have more resources to obtain grants or private donations that mitigate the loss of taxable property value. For example, Sims et al. (2022) find that median household income is strongly correlated with the amount of land protected in New England towns and cities since 1990.

The potential for fiscal impacts of land protection has also been considered within the context of preferential taxation of working lands. All U.S. states provide some form of tax relief to land used for forestry and agriculture, usually with the goals of promoting rural livelihoods and providing incentives to reduce the conversion of working lands to developed uses (Anderson 2012). The most common mechanism for tax relief is through current use provisions (also referred to as use-value assessment). Under current-use programs, eligible property is assessed for its income-producing capacity in agriculture and forestry, instead of its potential housing or commercial market value, which usually results in substantially lower tax obligations (Anderson and England 2015). Some early studies calculate the potential shift in property tax burden from agricultural to non-agricultural property under the assumption that lost property tax revenue is fully compensated for by increased tax rates on non-agricultural land, finding expected increases of 1–20% for
non-agricultural land (Chicoine and Hendricks 1985; Dunford and Marousek 1981; Ching and Frick 1970). Several county-level studies also estimate the foregone property tax revenue from current-use taxation by comparing property taxes paid under current-use and fair-market value, finding that current-use taxation may reduce property tax revenue by 35–75% (Coogan et al., 2014; Anderson and Grifﬁng 2000). However, a recent national study by Bigelow and Kueh (2022) using observational data on municipal ﬁnances found that while adoption of current-use taxation led to an 11% reduction in property tax revenue at the county level, the revenue loss was offset by increased transfers from state governments, leading to no overall impact on local revenues.

Cost of community services studies have also played an important role in public debates about land protection and tax rates (Clapp et al., 2018; Kotchen and Schulte 2009). These studies seek to compare the revenues gained from land in different uses to the costs of serving those land use types. These studies are accounting based and apportion municipal revenues and expenditures to speciﬁc land classes (e.g., residential development vs. commercial vs. open space) with the goal of comparing the ratio of expenditures to revenues for different land uses (Kotchen and Schulte 2009). The ﬁndings from this literature show that open space/farmland and commercial/industrial land uses often have expenditures to revenues ratios of less than one, meaning they “pay for themselves.” This is consistent with situations where most municipal expenditures are driven by needs such as schools, sanitation services and emergency services. Open space lands tend to “consume” few of these resources. However, the conclusions that can be drawn from this literature are limited because the ﬁndings often depend on the underlying assumptions of how to apportion budgetary costs to different land classes.

Our work is most closely related to the small number of empirical studies that examine the impact of permanent land protection on property tax rates or the property tax base. We are aware of only two: King and Anderson (2004) investigated the effect of new land protection with conservations easements in 29 Vermont towns, during a 10-year time period. They ﬁnd that property tax rates increase for up to four years after land is protected but that the impact of easement protection becomes insigniﬁcant or negative in the longer term. Vandegrift and Lahr (2011) examined the impact of contemporary and historical open space expenditures on property tax base growth in 566 New Jersey municipalities across a ﬁve-year period. They ﬁnd a small negative impact of contemporary land protection on tax base growth and no long-term impacts associated with historical expenditures on open space.

Our paper advances this literature by estimating the effects of land protection on tax rates using plausibly exogenous changes in land protection across a wide set of municipality types. By using panel data on more than 1400 towns and cities over a 20-plus year time span, we are able to isolate impacts using variation within municipalities over time and control for potentially confounding trends at the regional level and within time periods.

Studying New England also allows us to assess the impacts of both public and private land protection types. Private land protection, mainly through conservation easements, is playing an increasingly important role in preserving ecosystems and biodiversity across the country (Capano et al., 2019; Parker and Thurman 2019; Trust Alliance, 2015), yet is understudied relative to public land protection. In addition, we are able to assess heterogeneity in property tax impacts of land protection across towns and cities with an array of local economic characteristics. An understanding of impact heterogeneity is particularly important given growing concerns about equity in the beneﬁts and costs of environmental policies (e.g., Currie et al., 2021; Shapiro and Walker 2021; Carley and Konisky 2020; Colmer et al., 2020; Banzhaf et al., 2019). Significant disparities in access to protected land have been documented by factors including income and race (e.g. Sims et al., 2022; The Trust for Public Land 2020; Jennings et al. 2012). While greater equity in access to the beneﬁts of open space may be achieved partly through additional protection in disadvantaged communities, it is also important to understand the potential ﬁscal risks of new land protection across different community types.

3. Study area and data

To examine the impact of land protection on property tax rates, we assemble an annual panel at the municipal level. We combine data on municipal ﬁscal outcomes, land protection, and current use tax breaks, and socio-economic characteristics from 1990 to 2015 for 1436 municipalities (also “towns and cities”) in the New England region (Fig. 1, Table 1). We give a brief overview below, with further details of the region and data collection described in Appendix B.

3.1. Regional structure and property tax process

We examine municipal data from the ﬁve states of Massachusetts, Connecticut, New Hampshire, Vermont, and Maine (which comprise ﬁve of the six New England states and the large majority of land area). Municipalities in our study area are generally charged with providing local public goods and services such as schools and ﬁre-departments, collecting taxes to pay for them, and allowing building permits for new development. New England has weak county government structures and land is generally incorporated into municipalities (exceptions discussed in Appendix B). The population included in our study region was 13.6 million in 2015, and includes a continuum of urban areas like Boston, Hartford, and Worcester, dense to sparse suburban areas, and rural areas (U.S. Census Bureau, 2016).
Municipalities throughout New England have a similar process for setting municipal budgets. Generally, a budget or finance committee works in collaboration with municipal departments to prepare a budget for the upcoming fiscal year (Byrnes 2017; Massachusetts Municipal Association, 2014; Neal 2012; Vermont League of Cities and Towns, 2002; Hill 1992). The budget balances

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Fig. 1. Study area – land protection and land cover by 2015.
Notes: Land protected by 2015 as a percentage of area within a municipality from the Harvard Forest/Highstead database. Subfigure A shows total land protected, while Subfigures B-D show protection for mutually exclusive categories of protection: (B) land owned in fee by NGOs (C) land owned in fee by municipalities (D) easements on private land (E) land owned in fee by state/federal government. Figure F shows land cover in 2016 from the National Land Cover Database (Dewitz, 2019). High and Low density indicate developed land classes (with the “High” category here including both High and Intermediate density from the NLCD). Excluded areas are described in Appendix B.
Table 1
Summary statistics for differenced and level variables.

<table>
<thead>
<tr>
<th>Fiscal Variables (Regional Sample)</th>
<th>Obs.</th>
<th>Towns</th>
<th>Time Period</th>
<th>Units</th>
<th>Level Variables</th>
<th>Mean</th>
<th>SD</th>
<th>Units</th>
<th>Differenced Variables</th>
<th>Mean</th>
<th>SD</th>
</tr>
</thead>
<tbody>
<tr>
<td>Tax rate, equalized</td>
<td>9581</td>
<td>1436</td>
<td>1994-2015</td>
<td>$/$1000 value</td>
<td>11.17</td>
<td>5.61</td>
<td>$/$1000 value</td>
<td>0.0690</td>
<td>0.6318</td>
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<td></td>
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<tr>
<td>Tax rate, nominal</td>
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<td>1436</td>
<td>1994-2015</td>
<td>$/$1000 value</td>
<td>13.04</td>
<td>7.55</td>
<td>$/$1000 value</td>
<td>0.0705</td>
<td>0.8177</td>
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<td>Ln equalized value</td>
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<td>1436</td>
<td>1994-2015</td>
<td>Ln(dollars)</td>
<td>19.74</td>
<td>1.59</td>
<td>ΔLn (dollars)</td>
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<td>0.0551</td>
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<td>Ln assessed value</td>
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<td>Ln(dollars)</td>
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<td>1.58</td>
<td>ΔLn (dollars)</td>
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<td>0.0642</td>
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<td>Ln property tax levy</td>
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<td>1994-2015</td>
<td>Ln(dollars)</td>
<td>15.09</td>
<td>1.79</td>
<td>ΔLn (dollars)</td>
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<tr>
<td>Ln municipal revenues (CT,MA)</td>
<td>9581</td>
<td>519</td>
<td>1994-2015</td>
<td>Ln(dollars)</td>
<td>17.21</td>
<td>1.34</td>
<td>ΔLn (dollars)</td>
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<td>0.0242</td>
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<tr>
<td>Ln municipal expenditures (CT,MA)</td>
<td>9581</td>
<td>519</td>
<td>1994-2015</td>
<td>Ln(dollars)</td>
<td>17.10</td>
<td>1.34</td>
<td>ΔLn (dollars)</td>
<td>0.0200</td>
<td>0.0275</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

| Land Protection                   |      |       |             |       |                |      |    |       |                        |      |    |
| Total land protected as of prior time period | 9581 | 1436  | 1991-2012   | % town area | 15.68 | 15.12 | % town area | 0.2325 | 0.5033 |
| Ngo protection as of prior time period   | 9581 | 1436  | 1991-2012   | % town area | 1.80  | 3.48  | % town area | 0.0347 | 0.1234 |
| Municipal protection as of prior time period | 9581 | 1436  | 1991-2012   | % town area | 3.27  | 5.09  | % town area | 0.0230 | 0.0854 |
| Easement protection as of prior time period | 9581 | 1436  | 1991-2012   | % town area | 2.80  | 5.35  | % town area | 0.0987 | 0.2813 |
| State/federal protection as of prior time period | 9581 | 1436  | 1991-2012   | % town area | 7.55  | 13.17 | % town area | 0.0445 | 0.1841 |

| Current Use Value                 |      |       |             |       |                |      |    |       |                        |      |    |
| Land share in current use in 2010 (No MA) | 7140 | 1086  | 2010        | % town area | 35.90 | 22.15 | –       | –                       | –     | –  |
| Value share of land in current use in 2010 (MA) | 2447 | 350   | 2010        | % taxable value | 0.17  | 0.24  | –       | –                       | –     | –  |

| Socioeconomic Variables           |      |       |             |       |                |      |    |       |                        |      |    |
| Unemployment rate, prior time period | 9581 | 1436  | 1991-2012   | percent | 5.34  | 2.94  | percent | –0.0177 | 0.8099 |
| Labor force, prior time period     | 9581 | 1436  | 1991-2012   | labor force/acre | 0.43  | 1.15  | labor force/acre | 0.0010 | 0.0091 |

| Municipal Characteristics          |      |       |             |       |                |      |    |       |                        |      |    |
| Municipal area                     | 1436 | 1436  | 2010        | acres  | 19927.72 | 10647.26 | –       | –                       | –     | –  |
| Median household income            | 1436 | 1436  | 1990        | USD, thousands | 62.50 | 21.20 | –       | –                       | –     | –  |
| Vacation home share                | 1436 | 1436  | 1990        | percent | 15.42 | 18.37 | –       | –                       | –     | –  |
| Urban municipality                 | 1436 | 1436  | 1990        | 0/1     | 0.13  | 0.34  | –       | –                       | –     | –  |
| Exurban municipality               | 1436 | 1436  | 1990        | 0/1     | 0.35  | 0.48  | –       | –                       | –     | –  |
| Rural municipality                 | 1436 | 1436  | 1990        | 0/1     | 0.52  | 0.50  | –       | –                       | –     | –  |

Notes: Summary statistics showing average values and standard deviation for level and differenced variables at the three-year analysis time period. Level variables represent average values within three-year time periods used in the analysis, while differenced variables represent annual differences averaged over three years.
proposed expenditures against expected revenues and is then adopted or rejected directly by town residents at a town meeting or by a city council or another representative body in larger municipalities (Byrnes 2017; Massachusetts Municipal Association, 2014; Neal 2012; Vermont League of Cities and Towns, 2002; Hill 1992). Once a budget is approved, a property tax rate is set to raise the revenue required to cover the approved municipal appropriations in excess of other local revenues and transfers from the state (Reid 2012). The tax rate is set based on the value of taxable property in the municipality, according to the most recent valuation from the municipal assessor.

Municipalities in New England are quite dependent on property tax revenue, making it a good study area to detect impacts of land protection on property tax rates. The share of local government revenue from property taxes in 2015 was 55% in Connecticut, 54% in Maine, and 60% in New Hampshire, which is more than twice the national average of 27% (Urban Institute, 2020). In Massachusetts, the property tax revenue share was 44%, also considerably above average (Urban Institute, 2020). Other local revenue sources are relatively small: across the five states, in 2015, non-property tax revenues (own source) accounted for 14–27% of local revenues, relative to an average of 40% for local governments in the rest of the U.S. (Urban Institute, 2020). Sales taxes in our study region are primarily set at the state level. In Maine, New Hampshire, Massachusetts and Connecticut, municipalities cannot impose additional local sales taxes (Fritts 2021). In Vermont, local sales taxes are limited to 1% and only apply to a minority of towns (Department of Taxes 2022; Fritts 2021). None of the New England states allow local personal income taxes (Walczak 2019). Changes in fees for water or other services must be justified according to the expense of providing the service and so are unlikely to be affected by other needs for revenue (Division of Local Services 2016; Sanderson 2010). Given these constraints, options for municipalities to raise taxes or fees to make up for lost revenue from property taxes beyond raising the property tax rate are limited.

3.2. Municipal fiscal data and tax rate measures

Municipal fiscal data were obtained from each state’s department of revenue. Our main outcome of interest is the equalized property tax rate (also known as the effective property tax rate). The equalized property tax rate is calculated as the property tax levy (revenues needed from property taxes) divided by the full, fair market value of properties (the equalized property tax base; further explanation in Appendix B). This measure allows a more equitable comparison of effective tax burdens across jurisdictions because it is less affected by reappraisal cycles and state or local idiosyncrasies in assessment practices (Lincoln Institute of Land Policy and Minnesota Center for Fiscal Excellence, 2015; Bell and Kirschner 2009; Clapp et al., 2008). As a robustness check, we also consider the nominal tax rate, which is the rate that is published and depends on the assessed values of property (see Appendix B). Fig. 2 illustrates the spatial variation in property tax rates (A) and tax bases (B) within our study region. Differences in property tax systems across states create some challenges in constructing comparable measures across the region. To address these differences, we focus on property tax rates that are residential (if there are multiple rates) and are clearly controlled by the municipalities themselves (see Appendix B for further details). Finally, we also assemble data on the municipal tax levy, assessed and equalized taxable property values, as well as budget revenues and expenditures for the two states where these data are available (CT and MA). Fiscal variables for our study municipalities are summarized in Table 1 with more detailed definitions included in Appendix Table A1.

3.3. Land protection data and types

Land protection data are from the protected open space (POS) database (Harvard Forest 2020), which includes parcel-level spatial data on land ownership type, method of land protection (fee/easement), and the timing of the protection events. These data aggregate multiple data sources to provide a comprehensive layer of spatial land protection with a consistent schema of attributes (see Appendix B for further details). We measure land protection as a percentage of town land area to account for differences in town size.

Our analysis considers the change in four types of land protection: ownership by NGOs, ownership by municipalities, private conservation easements, and ownership by state or federal agencies. Conservation easements are voluntary, legal agreements between a landowner and a qualified NGO or public agency that extinguishes the right to develop the land. We restrict this category to easements on private land to create mutually exclusive categories for analysis. In the few cases where there is an easement on land in municipal, state/federal or NGO ownership, we characterize protection according to the ownership type.

The region has experienced large increases in land protection since 1990, as summarized in Fig. 3 and Fig. A1. In 1990, 12.8% of our study area was protected. Since then, all categories of land protection have increased, with the greatest increase from new easements on private land (Meyer et al., 2014), particularly in New Hampshire and Vermont (Fig. 3). Among land protected in our study area between 1990 and 2015, 52.3% of the new acreage was under private ownership with easement protection, 22.9% was under state/federal ownership, 14.8% was acquired in fee by NGOs, and 8.9% was protected through municipal acquisition. This reflects in part the substantial private land ownership in New England. Private landowners, including hundreds of thousands of family forest owners with small to mid-sized parcels, own more than 75% of the land across the region as a whole (Butler et al., 2016).

By 2015, 20.2% of our study area was protected, with 9.6% of the study area protected by state/federal government ownership, 2.2% through NGO ownership, 2.9% through municipal ownership, 5.1% through acquisition of easements on private land and 0.4% unclassified. Fig. 1 maps the percent of total land protected and by each protection type at the municipal level in 2015. Table 2 summarizes the four protection types and the potential parcel-level change in property tax obligations expected for each type.

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4 These data specifically capture outcomes at the municipal level, omitting property taxes levied separately by sub-municipal special districts like village, water-, fire-, and lighting-districts (see Appendix B).
Fig. 2. Municipal property tax rate and property tax base.
Notes: (A) Within state municipal percentile rank based on average equalized property tax rate from 1994 to 2015. The average is taken using the three-year analysis time periods used in estimation. Higher rank municipalities have higher property tax rates on average relative to other municipalities in the same state. (B) Within state municipal percentile rank of average equalized value per acre at baseline. The baseline years are 1992–1994 for municipalities outside of Vermont and 1996–1997 for Vermont municipalities.

Fig. 3. State and regional area protected over time, by land protection type.
Notes: Fig. 3 shows the percent of state and regional area protected in 1990 within our study area, and the change in percent area protected between 1990 and 2015, by land protection type and in total.
For fee simple acquisition of land by NGOs, the full taxable value of protected land would typically be removed from the property tax rolls. Additionally, in some cases, municipalities must raise the funds to purchase the land, which could require a tax increase or additional local fees. In terms of Equation (1), this means a reduction in the tax base as well as a potential increase in expenditures to fund the land acquisition. An increase in expenditures is not a given, however, as the land may be donated to the municipality or acquisition funds may come from state or federal grants or out of existing planned expenditures.

Conservation easements on private land usually result in a reduction in taxable value but not a full removal from the tax rolls. Taxable value is generally reduced because easement terms restrict future development and uses. At the same time, if this land was already enrolled in current use, if development was already restricted (e.g. by wetlands laws), or the land does not have high market value, then the impacts on the tax base may be limited.

Finally, we consider land acquired in fee by state and federal agencies, which also becomes tax-exempt. The federal government and all states in our sample except for Maine do make payments in lieu of taxes (PILOTs) to municipalities as compensation for the lost property tax revenue. Compensation varies and can be based on average tax rates in the state (DeNucci 2001), a fixed proportion of lost property tax revenue due to current use valuation (Vermont Department of Forests, Parks and Recreation 2010). If an easement decreases the market value of property in current use, that will reduce state payments to the town. Maine compensates municipalities 90% of lost property tax revenue due to current use, but only for losses related to the forestry focused Tree Growth program (Maine Revenue Services, 2020), which accounts for about 90% land in current use in Maine (Maine Revenue Services, 2010). Easements don’t affect compensation to towns for land in Tree Growth current use because payments are based on the difference between current use value and value of undeveloped land in the region, not the specific parcel that an easement is placed on (Maine Revenue Services, 2020). In Massachusetts, Connecticut, and New Hampshire, municipalities carry the cost of lost property tax revenue due to current use without compensation from the state.

### 3.4. Current use value assessments

As highlighted in the section above and Table 2, the expected change in the property tax base due to permanent land protection may also be related to current-use-value assessment programs. If land is already given tax breaks under current-use provisions at the time of protection, we expect the additional impact of permanent land protection on taxes to be smaller (Table 2). Substantial amounts of land were enrolled in current use provisions during our study period. State land shares in current use as of 2010 are: 30% in CT, 35% in ME (Maine Revenue Services, 2010), 38% in VT (Division of Property Valuation and Review 2011) and 52% in NH (Department of

---

**Table 2**

<table>
<thead>
<tr>
<th>Protection Type</th>
<th>Taxation Regime</th>
<th>From Tax Base</th>
<th>Payments that Offset Lost Tax Base</th>
</tr>
</thead>
<tbody>
<tr>
<td>1) NGO Fee Acquisition</td>
<td>Market Value Assessment</td>
<td>Total Taxable Value</td>
<td>Sometimes*</td>
</tr>
<tr>
<td>2) Municipal Fee Acquisition</td>
<td>Market Value Assessment</td>
<td>Total Taxable Value</td>
<td>No</td>
</tr>
<tr>
<td>3) Easement on Private Land</td>
<td>Market Value Assessment</td>
<td>Partial Taxable Value</td>
<td>No</td>
</tr>
<tr>
<td>4) State/Federal Fee Acquisition</td>
<td>Market Value Assessment</td>
<td>Total Taxable Value</td>
<td>Yes†</td>
</tr>
</tbody>
</table>

Notes: This table describes the expected first order fiscal impacts associated with each land protection type and pre-protection taxation regime of the land.

* Sometimes land trusts choose to make voluntary payments in lieu of taxes to municipalities to offset the property tax revenue loss resulting from their land acquisition. They are not required to do so however and there is not a systematic way of knowing who is making such contributions.

† The federal government and all states except for Maine make compensating payments in lieu of taxes to municipalities for state/federal owned land.

‡ Vermont compensates municipalities for lost property tax revenue due to current use valuation. Towns are compensated for the proportion of property tax revenue lost due to the difference in market and current use valuation (Vermont Department of Forests, Parks and Recreation 2010). If an easement decreases the market value of property in current use, that will reduce state payments to the town. Maine compensates municipalities 90% of lost property tax revenue due to current use, but only for losses related to the forestry focused Tree Growth program (Maine Revenue Services, 2020), which accounts for about 90% land in current use in Maine (Maine Revenue Services, 2010). Easements don’t affect compensation to towns for land in Tree Growth current use because payments are based on the difference between current use value and value of undeveloped land in the region, not the specific parcel that an easement is placed on (Maine Revenue Services, 2020). In Massachusetts, Connecticut, and New Hampshire, municipalities carry the cost of lost property tax revenue due to current use without compensation from the state.

---

5 Data on current use acreage for 2010 from the Forestry Division of the Connecticut Department of Energy and Environmental Protection.
Since the choice of a three-year time step is a somewhat arbitrary modeling decision, we include robustness checks with two-year and four-year average differences (Section 5.7). We model the change in tax rates as a function of the lagged change in land protection to reduce the possibility of reverse causality and to allow time for municipalities to adjust their tax rate following changes in land protection dates.

3.5. Municipal characteristics

We expect that the characteristics of a municipality are potential mediators of property tax impacts when new land protection occurs. To test these relationships, we collect data on the size and growth of each municipality’s property tax base, community type by residential density, the existing share of land protected, the share of vacation homes, and the median household income. Summary statistics are given in Table 1 and characteristics are mapped in Fig. A2. Details of the construction of each of these variables are described in Appendix B.

4. Empirical strategy

4.1. Main estimation approach


We use differences in tax rates and differences in land protection (rather than levels) to subtract out the time-invariant determinants of tax rates within each town. We average these differences across three years to reduce the influence of outliers in individual annual periods and to account for the fact that land protection dates may be measured with some lag because of event timing.

Since the choice of a three-year time step is a somewhat arbitrary modeling decision, we include robustness checks with two-year and four-year average differences (Section 5.7). We model the change in tax rates as a function of the lagged changes in land protection to reduce the possibility of reverse causality and to allow time for municipalities to adjust their tax rate following changes in land protection.

Our main regression model is therefore:

\[
\text{Ihs} \Delta \text{TaxRate}_{ic,t} = \beta \text{Ihs} \Delta \% \text{ProtLand}_{ic,t-1} + \text{Ihs} \Delta X_{ic,t-1} + \gamma_t + \Omega (t \times \lambda_t) + \epsilon_{ic,t}
\]

where \(i\) denotes each municipality, \(c\) is the metro-region of that municipality and \(t\) indexes each three-year time period. The dependent variable, \(\text{Ihs} \Delta \text{TaxRate}_{ic,t}\), is the Ihs transformed average difference in the equalized property tax rate. The policy variable of interest, \(\text{Ihs} \Delta \% \text{ProtLand}_{ic,t-1}\), is the Ihs transformed average difference in the percent of land protected. \(\text{Ihs} \Delta X_{ic,t-1}\) is a vector of controls for local economic conditions including: the Ihs transformed lagged average changes in the labor force, Ihs transformed lagged average changes in the unemployment rate, and the lagged average percentage change in the property tax base (tax base growth). While our use of changes-on-changes controls for time-constant municipal characteristics, we include this set of lagged time-varying controls to mitigate possible concerns that new land protection may be correlated with local economic conditions or appreciation of the tax base. In addition, we include \(\gamma_t\), a state-by-time-period fixed effect, and \((t \times \lambda_t)\) which controls for linear time trends for each metro-region. These account for possible differential trends in tax rates within sub-regions of New England. Our identifying assumption is that conditional on these controls, the remaining variation in new land protection is plausibly exogenous because it is driven by the uncoordinated activities of thousands of landowners and hundreds of land trusts, local governments, state, and federal agencies, as well as randomness in the timing of property transitions. We include the count of town observations by time-period and state used in the estimation in Appendix Table A2.

Our rationale for using the Ihs transformed differences—rather than the more standard approach of a difference in logged tax rates—is to reduce the influence of outliers in individual annual periods. We average these differences across three years to account for the fact that land protection dates may be measured with some lag because of event timing.

We use differences in tax rates and differences in land protection (rather than levels) to subtract out the time-invariant determinants of tax rates within each town. We average these differences across three years to reduce the influence of outliers in individual annual periods and to account for the fact that land protection dates may be measured with some lag because of event timing.

Since the choice of a three-year time step is a somewhat arbitrary modeling decision, we include robustness checks with two-year and four-year average differences (Section 5.7). We model the change in tax rates as a function of the lagged changes in land protection to reduce the possibility of reverse causality and to allow time for municipalities to adjust their tax rate following changes in land protection.

Our main regression model is therefore:

\[
\text{Ihs} \Delta \text{TaxRate}_{ic,t} = \beta \text{Ihs} \Delta \% \text{ProtLand}_{ic,t-1} + \text{Ihs} \Delta X_{ic,t-1} + \gamma_t + \Omega (t \times \lambda_t) + \epsilon_{ic,t}
\]

where \(i\) denotes each municipality, \(c\) is the metro-region of that municipality and \(t\) indexes each three-year time period. The dependent variable, \(\text{Ihs} \Delta \text{TaxRate}_{ic,t}\), is the Ihs transformed average difference in the equalized property tax rate. The policy variable of interest, \(\text{Ihs} \Delta \% \text{ProtLand}_{ic,t-1}\), is the Ihs transformed average difference in the percent of land protected. \(\text{Ihs} \Delta X_{ic,t-1}\) is a vector of controls for local economic conditions including: the Ihs transformed lagged average changes in the labor force, Ihs transformed lagged average changes in the unemployment rate, and the lagged average percentage change in the property tax base (tax base growth). While our use of changes-on-changes controls for time-constant municipal characteristics, we include this set of lagged time-varying controls to mitigate possible concerns that new land protection may be correlated with local economic conditions or appreciation of the tax base. In addition, we include \(\gamma_t\), a state-by-time-period fixed effect, and \((t \times \lambda_t)\) which controls for linear time trends for each metro-region. These account for possible differential trends in tax rates within sub-regions of New England. Our identifying assumption is that conditional on these controls, the remaining variation in new land protection is plausibly exogenous because it is driven by the uncoordinated activities of thousands of landowners and hundreds of land trusts, local governments, state, and federal agencies, as well as randomness in the timing of property transitions. We include the count of town observations by time-period and state used in the estimation in Appendix Table A2.

Our rationale for using the Ihs transformed differences—rather than the more standard approach of a difference in logged tax rates—is to reduce the influence of outliers in individual annual periods. We average these differences across three years to account for the fact that land protection dates may be measured with some lag because of event timing.
values—is based on its better performance in reducing the influence of outliers in our dataset as well as preserving the rank order of the within-municipality changes in the key variables (visualized in Appendix Fig. A3). Despite averaging across three years, our data does include some large changes in property tax rates and land protection. To ensure that potential outliers are not driving the results, we transform the average differences using the inverse hyperbolic sine (Ihs—see Burbridge et al. (1988)) and winsorize the furthest 1% of the data.\(^{13}\) In addition, we include robustness checks using the differences of log transformed variables and other possible specifications (Appendix Fig. A3, Section 5.7).

### 4.2. Heterogeneity in impacts, and impacts across time

To understand possible heterogeneity in the impacts of new land protection, we test for differential impacts by protection type and by municipal characteristics, including measures of land in current use, property tax base size and growth, municipality type based on housing density (rural, exurban, urban), percent of land protected in a municipality at the beginning of a time-period, share of vacation homes in municipal housing stock, median household income, and the lagged tax rate.

We first estimate a series of single interaction models with each characteristic included alone. However, as there may be important co-variation between these municipal characteristics, we also estimate two multiple interaction models. For these models we leave out housing density and land enrolled in current use due to data limitations and potential multicollinearity.\(^{14}\) We estimate the fully interacted model for aggregated changes in land protection and then for each protection type separately.

In addition, we test for the possibility that towns may adjust to land protection through a reduction in property tax revenue and expenditures on public goods. We have expenditures and revenue data from two states to test this directly. We also assess potential impacts on the total property tax levy and property tax base using data from all states. For these outcomes, there is an extremely wide range of values for the differences (from hundreds of dollars to hundreds of million dollars) which may make results sensitive to the choice of a scaling factor for the Ihs transformation.\(^{15}\) Therefore, for these dependent variables, we use the standard specification of difference in logs, keeping the right-hand side of our specification identical to Equation (2):

$$\Delta Ln(Y)_{t,ic} = \beta_{1} Ihs \Delta ProtLand_{e,j-1} + Ihs \Delta X_{t,ic} + \gamma_{o} + \Omega (t \times \lambda_{t}) + e_{ct}$$

(3)

where $$\Delta Ln(Y)_{t,ic} = Ln(Y)_{t,ic} - Ln(Y)_{t,ic-1}$$ and $$Y$$ represents our fiscal outcomes, including municipal expenditures, revenues, the property tax levy, and property tax base values. This specification models the relative (i.e. approximate percentage) changes in these fiscal variables as a function of the same transformed differences in land protection.

Finally, to assess the potential impacts of land protection on taxes over the longer term, we modify Equation (2) by introducing up to three time-period lags for the change in total land protection:

$$Ihs \Delta Tax \ Rate_{t,ic} = \sum_{j=1}^{3} (\beta_{j} Ihs \Delta ProtLand_{e,j-1}) + Ihs \Delta X_{t,ic} + \gamma_{o} + \Omega (t \times \lambda_{t}) + e_{ct}$$

(4)

where the subscript $$j$$ represents the temporal lag. Each lag represents a three-year time period, so this tests for impacts of land protection initiated up to nine years prior. We estimate this model to test the idea that tax rates might initially increase and then go back down again or that they may increase initially but not change further in the longer term.

### 5. Results

#### 5.1. Average impacts of land protection

Table 3 presents the results from our main specification, building up to this by including control variables in a stepwise fashion. This and subsequent figures and tables, unless noted otherwise, report estimated slopes evaluated at the mean values of the change in the tax rate and the lagged change in land protection. These slope estimates are derived by back transforming from the model coefficients and can be interpreted as the predicted change in the average annual equalized tax rate change ($ per $1000 value) resulting from a one percentage point increase in the average annual municipal area share newly protected in the prior three-year time-period. For reference, for the average municipality in our study area, one percent of town area is equal to 199.3 acres (Table 1). We explain the back transformation steps and provide an example in Appendix B.

---

\(^{13}\) We winsorize the top and bottom 1% of the distribution for all non-percentile variables used in estimation except for the change in land protection variables, for which we winsorize only the top 1%. Variables are winsorized after averaging to three-year time periods. We check robustness of our results to using non-winsorized data (Section 5.7).

\(^{14}\) Based on a review of correlations between variables (shown in Appendix Table A3), we exclude housing density from the fully interacted model due to the high degree of correlation between it and municipal property tax base size (>0.7). Town percentile rank by land enrollment in current use assessment is also excluded due to the relatively high correlation with property tax base percentile (>0.6) and because data for this variable is available for only one year (2010).

\(^{15}\) As illustrated in Bellemare and Wichman (2020) as well as Alhounton and Henningsen (2021), model results can be sensitive to the choice of a scaling factor for Ihs when there is a great range in the values. As a robustness check, we also estimate Equation (3) for these fiscal outcomes with both sides specified as differences of logs (Section 5.7).
Table 3
Estimated changes in municipal equalized tax rates as a function of lagged changes in land protected.

<table>
<thead>
<tr>
<th>Estimates</th>
<th>Δ TaxRate(_t)</th>
<th>Δ TaxRate(_t)</th>
<th>Δ TaxRate(_t)</th>
<th>Δ TaxRate(_t)</th>
<th>Δ TaxRate(_t)</th>
<th>Δ TaxRate(_t)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
<td>(4)</td>
<td>(5)</td>
<td>(6)</td>
</tr>
<tr>
<td>Δ % Protected (_t-1)</td>
<td>-0.0149 (^{*})</td>
<td>-0.0046 (^{*})</td>
<td>0.0318 (^{***})</td>
<td>0.0241 (^{**})</td>
<td>0.0241 (^{**})</td>
<td>0.0239 (^{**})</td>
</tr>
<tr>
<td></td>
<td>(0.0139)</td>
<td>(0.0142)</td>
<td>(0.0122)</td>
<td>(0.0110)</td>
<td>(0.0110)</td>
<td>(0.0110)</td>
</tr>
<tr>
<td>Δ Unemp. Rate (_t-1)</td>
<td></td>
<td>0.0036 (\hat{a})</td>
<td>0.0028 (\hat{a})</td>
<td>0.0028 (\hat{a})</td>
<td>-0.0018 (\hat{a})</td>
<td>-0.0018 (\hat{a})</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.0120)</td>
<td>(0.0120)</td>
<td>(0.0120)</td>
<td>(0.0119)</td>
<td>(0.0119)</td>
</tr>
<tr>
<td>Δ Labor Force/Acre (_t-1)</td>
<td></td>
<td></td>
<td>-2.9391 (^{***})</td>
<td>-2.9391 (^{***})</td>
<td>-3.0438 (^{***})</td>
<td>-3.0438 (^{***})</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>(0.4653)</td>
<td>(0.4653)</td>
<td>(0.4653)</td>
<td>(0.4653)</td>
</tr>
<tr>
<td>% Tax Base Growth (_t-1)</td>
<td></td>
<td></td>
<td></td>
<td>0.0126 (^{**})</td>
<td>0.0126 (^{**})</td>
<td>0.0126 (^{**})</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td>(0.0013)</td>
<td>(0.0013)</td>
<td>(0.0013)</td>
</tr>
</tbody>
</table>

Notes: Stepwise addition of controls. Column 1 indicates marginal effects from a regression of the lhs transformed change in tax rate on the lhs transformed lagged change in land protection without any controls. Column 2 adds CBSA time trends. Column 3 includes time-period fixed effects. In column 4, time-period fixed effects are replaced by state-by-time-period fixed effects. In columns 5 and 6, lhs transformed lagged average changes in labor force per acre and unemployment rate are added as controls. In column 7, we add lagged average tax base growth (%) as an additional control. Column 7 corresponds to the preferred main specification (Equation (2)). Table values are the estimated change in average annual equalized tax rate change ($/$1000 value) due to a one percentage point increase in prior period average annual municipal area share protected. Back transformation as described in Appendix B is used to calculate these values and the estimated standard errors. Standard errors in parentheses, clustered by municipality.

\( ^{*}p < 0.1; \quad ^{**}p < 0.05; \quad ^{***}p < 0.01 \)

We find that without controls, the relationship between the change in the tax rate and the change in prior period land protection is negative, with an estimated decrease in the average annual equalized tax rate change of $0.0149 per $1000 (Table 3, not statistically significant). With the addition of CBSA trends and state-by-time-period fixed effects, the estimated impact of new land protection becomes positive and significant with a magnitude of $0.0241 per $1000. The addition of further controls for labor market conditions and lagged property tax base growth does not substantially change the estimated impact, resulting in only a small reduction in the estimated magnitude of the tax rate change. Column 7 shows our preferred specification with the full set of controls as described in Equation (2). The estimated slope is $0.0231 per $1000 of value \((p = 0.036)\), again corresponding specifically to a one percentage point increase in the average annual municipal share of area newly protected in the prior time period (evaluated at average values of the change in tax rates and the change in land protection).

For comparison, we also represent this change in terms of an annual change in the property tax bill associated with a reference level of 100 acres of new land protection and for a typical home value (see Table 4). We estimate that 100 acres of new land protection results in an annual tax bill increase of $1.16 per $100,000 of value. Or, for an owner of a typical single-family home in New England, valued at $259,045, this translates to a tax bill increase of $3.00, which is a small change compared to the overall property tax bill (based on the average equalized tax rate across the region) of $2893 on that same home (Table 4).

100 acres is somewhat larger than the actual average change in protection among towns that did protect land: the average non-zero annual change is 84.9 acres and the average annual change overall is 45.8 acres. However, some towns saw several periods with larger increases. Therefore, for additional comparison, we calculate that the expected tax bill increase for the 90th percentile of annual non-zero increase in land protection (217.1 acres) would be $2.52 per $100,000 of value, or $6.53 for an owner of a typical New England home.

Small overall impacts on the property tax rate for the region could mask heterogeneity across individual states. We re-estimate the model separately for each state with results shown in Table 4. We find statistically significant impacts of land protection on tax rate change for some states, but all estimated slopes remain small (Table 4). An additional one percentage point increase in the average annual municipal area share protected results in a statistically significant increase in the property taxes in New Hampshire (estimate = $0.00546 per $1000 of value, \(p = 0.005\)) and a marginally significant increase in Vermont (estimate = $0.00292 per $1000 of value, \(p = 0.066\)). We do not find statistically significant impacts in Massachusetts, Connecticut, or Maine, with the estimated slopes indicating increases in Massachusetts and Connecticut that are similar to the overall estimated increase, and a possible small reduction in tax rates in Maine. Table 4 shows the estimated tax bill changes for each state, assuming a 100 acre per year increase in new protection. In comparison to the average tax bills expected based on typical home values and the municipal tax rates, each of these represents a small proportion of the average tax bill.

---

\( ^{16} \) $259,045 (2015 dollars) is the typical home value in our study area according to the Zillow Home Value Index (ZHVI) for 1996–2015. ZHVI reflects the mean value of single-family homes in the 35th to 65th percentile range of home values (Zillow Research 2019).

\( ^{17} \) Note that the average tax rates for NH and VT include only the municipal portion of those rates, not property taxes paid to the state to cover education; further explanation in Appendix B.
help to understand the magnitude of the tax bill change for the typical homeowner. 

$$0.0231 \times 100 \text{ acres} = 199.28 \text{ acres},$$

the adjustment factor is

$$\frac{100}{199.28}$$ 

municipal area share protected, for our New England study region and by state. These estimated slopes and standard errors are obtained by estimating Equation \((2)\) and back transforming as described in Appendix \(B\). Control variables include the prior period

year. The tax bill change is calculated by

1. adjusting estimated slopes to obtain the tax rate change per thousand dollars of taxable value for 100

2. multiplying the tax rate change (which is per $1000) by 100 to get tax bill change per $100,000 of value. Using column

3. standard errors are obtained by estimating Equation \((2)\) and back transforming as described in Appendix \(B\). Control variables include the prior period

4. time-period fixed effects, and linear CBSA specific time trends. Standard errors in parentheses, clustered by municipality. *p < 0.05; **p < 0.01.

Bottom section of table presents the corresponding estimated annual property tax bill change associated with 100 acres of new land protection per

While the average effect of new land protection on property tax rate change is small, including across states, this may still obscure important impact heterogeneity that is based on the type of land protection or municipal characteristics, rather than state. We next examine how the impact of land protection varies by type of protection and municipal characteristics using individual interactions and then fully interacted models that use aggregate change in protection and individual protection types (Figs. 4–6, Tables A4-A8).

### 5.2. Impact heterogeneity by land protection types

Table 4

<table>
<thead>
<tr>
<th>Estimates</th>
<th>New England</th>
<th>CT</th>
<th>MA</th>
<th>NH</th>
<th>VT</th>
<th>ME</th>
</tr>
</thead>
<tbody>
<tr>
<td>(\Delta \text{ % Protected}_{t+1})</td>
<td>0.0231**</td>
<td>0.0301</td>
<td>0.0266</td>
<td>0.0546***</td>
<td>0.0292*</td>
<td>−0.0513</td>
</tr>
<tr>
<td></td>
<td>(0.0110)</td>
<td>(0.0674)</td>
<td>(0.0200)</td>
<td>(0.0196)</td>
<td>(0.0159)</td>
<td>(0.0318)</td>
</tr>
<tr>
<td>(N)</td>
<td>9581</td>
<td>1010</td>
<td>2447</td>
<td>1577</td>
<td>1425</td>
<td>3122</td>
</tr>
<tr>
<td>(R^2_{\text{adj}})</td>
<td>0.4453</td>
<td>0.5223</td>
<td>0.5919</td>
<td>0.2837</td>
<td>0.2118</td>
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<td>Municipalities</td>
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<td>169</td>
<td>350</td>
<td>230</td>
<td>241</td>
<td>446</td>
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<tr>
<td>LaborMarketControls</td>
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<td>Yes</td>
<td>Yes</td>
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<td>Yes</td>
<td>Yes</td>
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<tr>
<td>State-By-Time-Period FE</td>
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<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
</tr>
<tr>
<td>Time-Period FE</td>
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<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>CBSATrends</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>TaxBaseGrowth</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>

Notes: Top section of table shows the estimated changes in average annual equalized tax rate change ($/$1000 value) due to a one percentage point increase in the prior period average annual municipal area share protected, for our New England study region and by state. These estimated slopes and standard errors are obtained by estimating Equation \((2)\) and back transforming as described in Appendix \(B\). Control variables include the prior period

1. average changes in labor force per acre and unemployment rate, average prior period tax base growth (\%), time-period or state-by-time-period fixed effects, and linear CBSA specific time trends. Standard errors in parentheses, clustered by municipality. *p < 0.1; **p < 0.05; ***p < 0.01.

Bottom section of table presents the corresponding estimated annual property tax bill change associated with 100 acres of new land protection per

year. The tax bill change is calculated by

1. adjusting estimated slopes to obtain the tax rate change per thousand dollars of taxable value for 100

2. multiplying the tax rate change (which is per $1000) by 100 to get tax bill change per $100,000 of value. Using column

3. standard errors are obtained by estimating Equation \((2)\) and back transforming as described in Appendix \(B\). Control variables include the prior period

4. time-period fixed effects, and linear CBSA specific time trends. Standard errors in parentheses, clustered by municipality. *p < 0.1; **p < 0.05; ***p < 0.01.

While there is considerable randomness in the overall process and timing of land protection, changes in protection may still reflect some differential selection, particularly among individual protection types (Fig. A1). From our visualization of the changes, we note that municipal protection is more common in the faster growing and more populous southern New England, while easements and state/federal land protection are more prevalent in more rural areas in western and northern New England. Some NGO protection in

in turn is occurring in amenity rich areas with more vacation homes (Fig. A1). Land protection types may therefore interact with other municipal characteristics. For this reason, we estimate models allowing for interaction between protection types and municipal characteristics. All estimated coefficients are evaluated at the mean values of the change in tax rates and the change in land protection and for mean values of other variables in the multiple interaction models.

Our results from the multiple interaction models indicate a similar pattern for the overall impacts of different land protection types,
with slightly larger estimated tax rate changes (Fig. 6A and Table A8). The largest increase in tax rate change is associated with municipal protection (estimate = $0.103, p = 0.045, evaluated at the means of all variables), followed by conservation easements (estimate = $0.0480, p = 0.042), state/federal protection (estimate = $0.0317, p = 0.26), and NGO protection (estimate = $0.0318, p = 0.41).

In terms of tax bill changes, assuming an increase in new protection of 100 acres, these results by type translate to annual tax bill increases of $5.17 (municipal protection), $2.41 (private easements), and $1.59 (state/federal protection) per $100,000 of value, while NGO protection is associated with a tax bill decrease of -$1.60 per $100,000 of property value. Our conservation easement results fall near the range of property tax impacts estimated by King and Anderson (2004), the most relevant previous study.\footnote{They estimate that protecting 100 acres using conservation easements can result in tax bill increase ranging from $8.40 to $15.80 per $100,000 of value in the three years following the change in land protection. Our lower estimates are consistent with our use of only the municipal portions of the property tax rate for Vermont and New Hampshire (King and Anderson use the aggregate property tax rate).}

These findings may be partly explained by how different protection types are expected to affect property taxes, as discussed in Section 3.3. In particular, municipal land acquisition may require raising funds to purchase land, so the expected impact could be greater than for other types. Easement protection may result in smaller changes in tax rates, compared to municipal land acquisition, because only a portion of taxable property value is lost when an easement is established and many lands were in current use assessment prior to being protected. The fiscal impacts of state and federal protection may be partially offset by PILOT payments. The possible

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**Fig. 4.** Heterogeneity in impacts of new land protection on tax rates, single interaction models.

Notes: Plots A-H show heterogeneity in the estimated impacts of new land protection on equalized tax rates, by protection type and municipal characteristics. The plots present the expected change in the average annual equalized tax rate change for a 1 percentage point increase in the annual municipal area share protected in the prior time-period. The expected changes are back transformed estimates based on Equation (2) with the addition of interaction terms; see also Appendix Table A4 for numerical results and Appendix B for explanation of back transformation. Tax rate change estimates are evaluated at 10th-90th percentile values of the variables used in the interaction, in the case of continuous variables, while holding other variables at their means. The 10th, 50th and 90th percentile values of the individual terms being evaluated are labeled on the X axes.
reduction in tax rates due to NGO protection is somewhat surprising. Although this result is not precisely estimated, if the true impact on tax rates is indeed negative, this could be explained by amenity effects or increased recreation spending associated with these properties that contributes to other municipal tax revenues. Anecdotally, NGOs in the region have invested considerably in infrastructure for recreation or historical visitation and worked in partnership with towns, possibly contributing to a positive impact on local property values.

5.3. Heterogeneity by municipal characteristics: current use, tax base size, tax base growth, density, tax rates

A possible explanation for the overall small impacts of land protection on taxes is that many lands were already paying relatively low taxes due to current use value assessment programs prior to acquisition or easements. Although data on current use enrollment is limited, we test this by interacting the percentile rank of land in current use with new land protection. We find that the tax rate impacts do decline with increasing amount of land in current use (Fig. 4B and Table A4, Column 2) and may even be negative for towns with large amounts of enrolled land.

The size and growth of the local property tax base as well as the existing tax rate may also play an important role in how tax rates respond to new protection. Smaller tax base towns may experience proportionally greater losses of taxable property value for the same dollar value change. At the same time, many towns with small tax bases are rural and face less development pressure, so parcel values...
Fig. 6. Heterogeneity in impact of new land protection on tax rates by land protection types, multiple interaction model.

Notes: Figures A–F show how the tax impacts of land protection vary by protection type and with respect to individual terms included in the multiple term interaction model. These figures present expected change in the average annual equalized tax rate change for a 1 percentage point increase in...
the annual municipal area share protected by a given type of land protection, in the prior time-period. See Appendix B for explanation of back transformation. Tax rate change estimates are evaluated at 10th-90th percentile values of the variables used in the interaction while holding other variables at their means. The 10th, 50th and 90th percentile values of the individual term being evaluated is labeled on the X axes. See Appendix Table A7 for the interaction model coefficients used to compute these changes and Table A8 for the numerical results plotted here.

may be less, meaning that dollar value changes are less for the same size and land use type.

In the single interaction model, we find that on average, the tax impacts of new land protection do not vary substantially by property tax base size (Fig. 4C and Table A4, Column 3). This result is somewhat surprising, since we might expect towns with large tax bases to be better positioned to absorb a loss of property tax revenue when land is protected. Instead, we find that the rate of tax base growth generally matters more than the level of the tax base (Fig. 4D, and Table A4, Column 4). The slope of the interaction term between new land protection and tax base growth is negative and statistically significant. This indicates (as also shown in Fig. 4D) that new land protection does result in significant increases to tax rates when it coincides with low rates of tax base growth.

The implication of this result is that the towns that may need to worry most about seeing larger tax increases from land protection are those that are experiencing slow growth in their tax base over time. However, as Fig. 4E (and Table A4, Column 5) shows, this cannot be easily predicted simply on the basis of whether areas are more urban vs. more rural. We find that the magnitude of the increase in property tax rate change is significantly different from zero only for exurban towns (estimate: $0.0309 per $1000, \( p = 0.049 \)). Although this is imprecisely estimated, towns and cities classified as urban had the largest estimated increase in tax rate change (Fig. 4E). Rural areas, which often draw the most attention in debates about land protection and might be expected to be growing more slowly, did not see large or statistically significant impacts on taxes on average as a result of land protection. We see similar results in Fig. A4, which gives a comparison of results based on using a continuous measure of housing density (see Appendix B for more details). Finally, we also find that the local tax rate level influences the impact of land protection. The interaction between the change in land protection and the prior period tax rate in the single interaction model is negative and statistically significant (\( p = 0.007 \)), suggesting that municipalities with higher lagged tax rates see smaller tax rate increases from land protection (Column 9, Table A4). It is possible that municipalities with already high rates are more reluctant to increase them further.

Interestingly, our results indicate that while additional tax base growth may reduce the tax rate impacts of land protection, tax base growth by itself is consistently associated with an increase in property tax rates (as shown across all columns in Table A4, last row of the table). While speculative, this is consistent with the findings from the costs of community services literature suggesting that growth actually raises taxes because the high costs of new municipal services outweigh the additional revenue brought by new development (Murray and Catanzaro 2019; Clapp et al., 2018; Kotchen and Schulte 2009).

In the fully interacted model, we find generally similar results (Figs. 5 and 6). However, for tax base size (Fig. 5B, Table A6) we find suggestive evidence of higher tax impacts for smaller tax base towns. For towns at the 10th percentile of tax base size, the estimated increase in tax rate change is $0.0349 (\( p = 0.055 \)). For 100 acres of new protection, this translates to a tax bill increase of $1.75 per $100,000 of value. Estimates from the multiple interaction model by protection type suggest that this result may be driven by municipal land protection (Fig. 6B, Table A8). Note that the estimated increase in tax rate change from new municipal protection for towns at the 10th percentile of tax base size is about 8 times larger than the overall average (estimate = $0.190, p = 0.056). For 100 acres of new protection, this translates to an annual tax bill increase of $9.53 per $100,000 of value, which is among the highest impacts that we find.

In the fully interacted model, tax base growth again appears to be a more consistent predictor of the impacts of land protection on tax rates (Figs. 5C and 6C). We find that the increase in tax rate change associated with new land protection decreases with the rate of tax base growth. This finding is consistent across both of our multiple interaction models, for total change in land protection and all land protection types. We observe some of the highest increases in tax rate change in our analysis for the slowest growing towns (see 10th percentile results in Tables A6 and A8), with increases in tax rate change as high as $0.0580 per $1000 (\( p = 0.002 \)) for total change in land protection (Table A6) and increases of $0.148 per $1000 (\( p = 0.033 \)) for municipal protection and $0.0952 per $1000 (\( p = 0.002 \)) for easements (Table A8). For a 100 acre increase in new annual land protection, these increases in tax rate change represent annual tax bill increases of $2.91, $7.43 and $4.78 per $100,000 of value for total change in land protection and municipal and easement protection, respectively, holding all other variables at their means. This highlights the potential role of growth as an important factor for mitigating tax rate increases that can result from land protection, even though growth by itself does not necessarily reduce tax rates.

5.4. Municipal characteristics: existing land protection, second homes, median income

In addition to growth rates, debates about land protection and tax rates often focus on possible constraints to development (or at least development of single-family homes with large lot sizes) that are posed when a high share of town land is set aside as protected. For this reason, we test whether the impact of additional protection is different for municipalities with a high cumulative share of area already protected at the beginning of the prior three-year time period. Fig. 4F (and Table A4, Column 6) shows that on average, there is no significant variation in the impact of land protection by land share protected (land share is measured at the start of the prior period).
Fig. 6. (continued).
In fact, if anything, there are smaller expected increases in taxes in the towns with a high share of land already protected. These results may indicate either that most towns in the region had not yet hit “build-out” constraints, or that such concerns can be overcome by redevelopment and increased density of housing. Alternately, a high share of land protection may substantially raise the value of the existing housing stock. Regardless, this result is important to note as it runs against a conventional wisdom that taxes will generally increase the most where there is already a lot of protected land.

This result generally holds in the multiple interaction models, with some exceptions. Estimates from the multiple interaction model with total change in protection show no significant differences in tax impacts with share of land protected (Fig. 5C, Table A8). Results from the multiple interaction model with individual protection types actually indicate a pattern of lower expected tax impacts for towns that already have a high share protected for NGO, municipal and private easement protection (Fig. 6D, Table A8). For these protection types, the towns and cities most affected by tax increases appear to be those with little existing land protection. In particular, among municipalities in the 10th percentile of already protected land, we find increases in tax rate change of $0.2065 per $1000 (p = 0.031) for municipal protection and $0.0784 (p = 0.013) for easement protection. These increases in tax rate change are associated with tax bill increases of $10.36 and $3.93 per $100,000 of value for municipal and easement protection, assuming 100 acres of additional annual protection. However, for state and federal protection, we estimate higher increases in tax rate change with more land protected at baseline rather than less. The estimated increase in tax rate change is $0.0897 per $1000 (p = 0.024) for towns in the 90th percentile of pre-existing land protection (>34% land area protected). Among the towns in the 90th percentile of pre-existing land protection, state and federally owned land on average accounts for 73% of total protected land. The larger tax rate impacts associated with new state and federal protection in these towns may potentially reflect the cumulative impact of state and federal land ownership and associated PILOT payments that don’t fully offset the lost property tax revenue (Bump 2020; Pinho and Dilworth 2020; DeNucci 1994, 2001). This suggests that special attention is warranted to assess potential tax impacts in towns where new state and federal protection is proposed and large amounts of land are already in reserves.

Communities may also be better positioned to benefit from land protection if they have a larger number of second homes or greater average incomes. The benefits of land protection may capitalize into local property values faster and to a greater extent in scenic, high amenity areas as well as in higher income towns with greater ability to pay for amenities. Second homes also tend to use fewer local services and may be associated with areas particularly rich in natural amenities, such as near coastlines or lakes (Polyakov et al., 2013; Irwin et al., 2010). Consistent with these expectations, we find that where there is a larger share of vacation homes in the municipal housing stock, there are smaller impacts on property tax rates due to new land protection (Fig. 4G and Table A4, Column 7). These results also hold based on estimates from the multiple interaction model for total protection (Fig. 5D) and individual protection types (Fig. 6E).

In addition, we find that the impact of land protection on tax rates increases as municipal-level median household income decreases (Column 8, Table A4 and Fig. 4H). For low-income towns (10th percentile rank), we estimate an increase in the tax rate change of $0.0357 per $1000 (p = 0.049) compared to high income towns (90th percentile rank) where the estimated increase in the tax rate change is $0.0079 (p = 0.638). In the fully interacted models (Figs. 5E and 6F), we also find that income is a consistent predictor of tax rate change for total protection and all protection types, with the largest increase in tax rates resulting from new municipal land protection in low-income municipalities. Specifically, for towns at the 10th percentile of median household income, we find increases in tax rate change of $0.0525 per $1000 (p = 0.010) based on total change in protection (Appendix Table A6), and $0.1525 (p = 0.068) for municipal protection, $0.0797 (p = 0.021) for easements and $0.0854 (p = 0.079) for state/federal protection (Appendix Table A8). For a 100 acre increase in new annual protection, these increases in tax rate change represent annual tax bill increases of $2.63, $7.65, $4.00 and $4.29 per $100,000 of value for total, municipal, easement and state/federal protection. These results indicate that greater attention to the potential tax impacts of land protection is warranted for lower-income municipalities.

5.5. Impact of land protection on levies and expenditures

A fundamental concern about land protection is that it could affect spending on other public goods that municipalities provide. If land protection results in a loss of property tax revenue that is not made up for by increases in taxes, towns may be forced to reduce expenditures. However, using the data from Massachusetts and Connecticut where expenditure data is available, we do not find evidence consistent with these concerns. 19 As shown in Table 5, we find that the estimated impacts of new land protection on expenditures and revenues are actually positive, although only marginally significant for expenditures (p = 0.062) and not substantial in magnitude compared to the overall average expenditures ($43,610 for a 100 acre change, compared to an average of $26.70 million). 20

In addition, using data from both MA and CT as well as all five states, we test for impacts on the property tax levy and total property values (assessed and equalized). We find a small, marginally significant estimated increase in the property tax levy for MA and CT (Table 5, Column 3, p = 0.090). The increase is also positive for all five states although not statistically significant, and again the

19 We note here that the levy increase estimated for the Massachusetts and Connecticut sample may be affected by Proposition 2.5 in Massachusetts, which limits the year-to-year tax levy increase to 2.5% (Division of Local Services 2007). However, municipalities can also choose to override this limit by a majority vote of the electorate and city council; 38% of such requests in MA were successful between 1990 and 2007 (Roscoe 2014).

20 We use logged values of the municipal expenditures, revenues, levy, and assessed as well as equalized values in our analysis. To obtain means in dollars, we exponentiate the average logged values presented in Table 1. For expenditure, exponentiating 17.10, yields $26,695,351.
Table 5
Impact of land protection on fiscal outcomes.

<table>
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<tr>
<th></th>
<th>ΔLn Expenditure$_{t}$</th>
<th>ΔLn Revenues$_{t}$</th>
<th>ΔLn Levy$_{t}$</th>
<th>ΔLn AssdVal$_{t}$</th>
<th>ΔLn EqVal$_{t}$</th>
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<tr>
<td><strong>MA &amp; CT Only</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
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<tr>
<td>Δ % Protected$_{-1}$</td>
<td>0.0024*</td>
<td>0.0018</td>
<td>0.0016*</td>
<td>-0.0009</td>
<td>-0.0012</td>
</tr>
<tr>
<td>(0.0013)</td>
<td>(0.0012)</td>
<td>(0.0009)</td>
<td>(0.0017)</td>
<td>(0.0017)</td>
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</tr>
<tr>
<td>Change for 100 acre annual increase in protection</td>
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<td>$35,647</td>
<td>$19,298</td>
<td>$-636,934</td>
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<td>3457</td>
<td>3457</td>
<td>3457</td>
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<td>0.2746</td>
<td>0.2656</td>
<td>0.6627</td>
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<td><strong>All States</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Δ % Protected$_{-1}$</td>
<td>–</td>
<td>–</td>
<td>0.0021</td>
<td>-0.0023</td>
<td>-0.0011</td>
</tr>
<tr>
<td>(0.0014)</td>
<td>(0.0016)</td>
<td>(0.0009)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Change for 100 acre annual increase in protection</td>
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<td>–</td>
<td>$3812</td>
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</tr>
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<td>–</td>
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<td>9581</td>
<td>9581</td>
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<td>0.6741</td>
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<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>State-By-Time-Period FE</td>
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<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>CBSATrends</td>
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<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
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<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
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</table>

Notes: This table shows the estimated predicted change in the average annual log difference for a one percentage point increase in prior period average annual municipal area share protected. These estimated marginal effects, standard errors and the corresponding dollar changes are obtained by estimating Equation (3) and back transforming coefficients as described in Appendix B. Marginal effects in $ are evaluated at the mean values of fiscal outcomes and lagged land protection for a 100 acre increase in annual land protection.

The same control variables are used in all Table 5 regressions: lhs transformations of the prior period average changes in labor force per acre and unemployment rate, average prior period tax base growth (%), state-by-time-period fixed effects, and linear CBSA specific time trends. Change in revenues and expenditures are outcomes from municipal budgets and are available only for MA & CT, while changes in the municipal levy, assessed values and equalized values are available for all states. Standard errors are in parentheses, clustered by municipality. *p < 0.1; **p < 0.05; ***p < 0.01.

Table 6
Impact of land protection over time.

<table>
<thead>
<tr>
<th></th>
<th>ΔTaxRate$_{t}$</th>
<th>ΔTaxRate$_{t}$</th>
<th>ΔTaxRate$_{t}$</th>
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</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
</tr>
<tr>
<td><strong>3 Yr Lag: Δ % Protected$_{t-1}$</strong></td>
<td>0.0231**</td>
<td>0.0145</td>
<td>0.0151</td>
</tr>
<tr>
<td>(0.0110)</td>
<td>(0.0124)</td>
<td>(0.0137)</td>
<td></td>
</tr>
<tr>
<td><strong>6 Yr Lag: Δ % Protected$_{t-2}$</strong></td>
<td></td>
<td>-0.0129</td>
<td>-0.0089</td>
</tr>
<tr>
<td>(0.0118)</td>
<td>(0.0138)</td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>9 Yr Lag: Δ % Protected$_{t-3}$</strong></td>
<td></td>
<td></td>
<td>0.0006</td>
</tr>
<tr>
<td></td>
<td></td>
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<td>(0.0124)</td>
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<td>LaborMarketControls</td>
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<td>Yes</td>
<td>Yes</td>
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<tr>
<td>State-By-Time-Period FE</td>
<td>Yes</td>
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<td>CBSATrends</td>
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</tr>
<tr>
<td>TaxBaseGrowth</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>

Notes: Estimated slopes show the change in average annual equalized tax rate change ($/$1000 value) due to a one percentage point increase in prior period average annual municipal area share protected as well as in additional lagged periods. The 3-year lag is the standard lag used in our analysis and represents the average change in the prior time period. The 6- and 9-year lags represent average change in protection two and three time periods prior. These slopes and standard errors are obtained by estimating Equation (4) and then back transforming as described in Appendix B. Control variables include the lhs transformations of the prior period average changes in labor force per acre and unemployment rate, average prior period tax base growth (%), state-by-time-period fixed effects and linear CBSA specific time trends. Standard errors in parentheses, clustered by municipality. *p < 0.1; **p < 0.05; ***p < 0.01.
estimated dollar values of both are small ($19,206 and $3801).

Finally, we find decreases in total assessed and equalized property values, although they are not statistically significant for either MA and CT or our region as a whole (Table 5, columns 4 and 5). Here we would expect to see a negative first order impact as land protection takes land off the tax rolls, but this may be offset by a secondary effect of potential increases in amenity values. Given the negative coefficients on assessed/equalized values and the positive coefficients on levies as well as the tax rate, our results are consistent with a scenario where towns may slightly overshoot in adjusting tax rates. In other words, they may set tax rates based on expectations of lost revenue from land protection without counting on potential gains from amenity effects. Municipalities then end up raising slightly more in levies than expected due to amenity effects raising the value of surrounding properties.

Crucially, these results suggest that municipalities were not fiscally constrained by land protection. However, while the options to raise revenue from other sources are limited, non-property tax channels could potentially have been used to offset lost property tax revenue. Further exploration of these non-property tax channels of fiscal adjustment using different datasets is outside the scope of our study but would be a fruitful avenue for future research.

5.6. Impacts over time; amenity effects

To understand the temporal trajectory of the tax rate change, and to test for the possibility that tax rates may decrease in the long run due to tax base growth from the amenity effects associated with land protection (Chamblee et al., 2011; Anderson and West 2006; Geoghegan et al., 2003; Irwin 2002; Thorsnes 2002), we include additional lagged values of land protection and estimate Equation (3). Table 6 builds up the lagged model results, adding one lagged time period at a time up to three lags, which together span a nine-year time period. If tax rates were to go up and then back down again as a result of land protection, we should expect that the prior period coefficient (3-year lag) is positive, while the preceding lags (6 and 9-year lags) are negative, or become increasingly negative. If tax rates were to go up once but not increase further, then we should expect that the prior period coefficient is positive while the preceding ones are zero. Our lagged model estimates are not statistically significant but show a pattern of estimated positive impacts on tax rate change in the first lagged period, negative in the second and close to zero in the third. Together with our main estimates, we interpret these results as being consistent with a small tax rate increase due to recent land protection, that does not continue to increase and possibly may come back down in the intermediate term, without additional long term impacts.

The possibility that increased amenity values due to land protection may be capitalized into property values also raises the issue that our estimated tax bill change due to a municipal tax rate increase may be a lower bound. In a national study, Lang (2018) found that housing prices at the zip code level increased by 0.68–1.12% for every $1000 of open space spending per household authorized through open space referenda. Similarly, in a case study of two Massachusetts towns, Heintzelman (2010) showed that the passage of the Community Preservation Act to fund open space and historical preservation was associated with local property value growth of 1.5–4.5% (although these benefits accrued only to homes above the 65th percentile of value).

Applying Lang’s estimates to our region, we roughly estimate that an increase in protection of 100 acres may result in property value growth ranging from 0.14% to 0.43% at the regional level, leading to an increase in tax bills of $1.40–$7.47 per $100,000 of value. At the state level, property value growth range may range from 0.048% to 1.02%, with associated tax bill increases of $0.66 - $12.83 per $100,000 of value (see Appendix B for details of the calculations and Appendix Table A9 for estimates by state and region). These estimates indicate that the potential increase in the tax bill due to capitalization of amenity value is comparable in magnitude, and could possibly be even larger, than the increase due to tax rate change.

In addition, while our results indicate important heterogeneity by municipal characteristics, they cannot measure important possible heterogeneity at the parcel level, driven by very localized amenity effects. Future research at the parcel level would help to understand the specific incidence of costs and benefits associated with land protection at the household level.

5.7. Robustness checks

Our estimates for the overall impact of land protection are generally robust to alternative decisions about the length of time steps, treatment of outliers and model specification. Appendix Table A10 compares estimated slopes for the average effect, and by protection type, across alternative averaging time steps. Results from two- and three-year averaging periods look very similar, while impacts on tax rate change overall are estimated as close to zero for the four-year averaging period. This is consistent with potentially smaller long-term impacts over time or could possibly reflect a greater amenity effect over time associated with NGO owned land. Appendix Table A11 presents estimated slopes with and without winsorization and where outliers are dropped instead of winsorized. The overall results are robust to the treatment of outliers (Table A11).22

Table A12 (Columns 2–5) presents estimates based on alternative functional forms of the main specification. We include percentage differences, log differences, untransformed differences of the tax rates and land protection and an alternative model using municipal fixed effects rather than a differences approach. The magnitude of the estimated tax bill change for 100 acres of new protection varies across these models from $0.35 to $0.90 per $100,000 of value and indicates that our overall conclusion that new land protection has only small impacts on tax rates is not being driven by our choice of model specification. In Table A13, we provide a robustness check of

21 Robustness checks using annual data and successive annual lags confirmed these general patterns.

22 We prefer winsorization to dropping outliers to ensure that individual observations do not drive the results because we do not have a clear justification to drop them and we do not want to further unbalance the panel unless necessary.
results using the published tax rate, also called the nominal tax rate following conventions in the literature (Song and Zenou 2006; Dye et al., 2001; Mikesell 1980). Since this is the rate seen by local property owners, it may be most salient. However as noted in the main text, nominal tax rates do not account well for differences in assessment practices or timing across localities. We find similar average annual tax rate increases associated with a one percentage point annual increase in municipal area protected: $0.025/$1000 of value with nominal tax rate (vs $0.0231/$1000 of value with equalized tax rate).

Finally, as a robustness check of Table 5, we estimate the impact of land protection on revenues, expenditures, levy and property tax base where we specify both the change in the fiscal variables and land protection as log differences (Appendix Table A14). The signs of the estimates are consistent, with somewhat larger estimated magnitudes and a marginally significant decrease in assessed values for all states ($p = 0.091$) and a marginally significant increase in the property tax levy for all states ($p = 0.085$).

6. Conclusion

Public desire for additional protected open space and also for low property taxes is a perennial source of tension and debate. The aim of this study was to estimate the impacts of new land protection on property tax rates, and to test for heterogeneity in impacts across protection types, municipal characteristics and conditions that may amplify or moderate these effects. Using data from more than 1400 towns and cities in New England, we analyzed the impacts of new land protection, using a panel data estimation approach comparing changes in tax rates following changes in land protection within municipalities.

Our results indicate that on average, the tax impacts of new public and private land protection are small, adding just a few dollars to the annual tax bill for most homeowners in the short run. These results suggest that for the majority of towns and cities, new land protection can be achieved without substantial impacts on other taxpayers or on the provision of public goods. The local benefits of this protection, including recreational opportunities, preservation of cultural heritage, wildlife habitat, and ecosystem services such as improved water quality, decreased flood risk and increased climate resilience may be considerably larger in value to residents than the modest increase in property tax bills.

While the impacts are typically small, they are heterogeneous, with some towns and cities likely to experience relatively larger tax rate increases than others. This includes municipalities with slowly growing tax bases, fewer vacation homes, lower average household incomes and less land enrolled in current use taxation. We also found greater tax impacts for towns that engaged in substantial municipal protection when they had low growth rates or small tax bases, and for towns that received state and federal protection when they already had a very high share of land protected.

These results highlight disparities in impacts and suggest that the towns least able to afford increases in property tax rates may also be those most likely to experience the greatest impacts. In addition, while potential tax costs are borne by the municipalities where the protection occurs, the benefits of land protection often extend to communities throughout the region through visitation, improved air and water quality, climate mitigation and other ecosystem services. In recognition of these broader public benefits, state and federal agencies can support local fiscal health by ensuring that payments in lieu of taxes programs are fully funded and are large enough to provide real compensation for the value that these protected lands provide, particularly in communities with fewer fiscal resources. Public and private organizations can also play a role in ensuring access to additional funds for land protection. For example, current requirements for municipalities to provide matching funds or prepare open space plans in order to receive state or federal grants for municipal land protection may create barriers to accessing outside funds that can be overcome with additional public or private assistance. Agencies and NGOs engaged in land protection should be aware of how the likely fiscal impacts of land protection may depend on where this protection occurs and can be proactive in sharing strategies or resources that support healthy municipal budgets and empower local communities to make sustainable decisions about new land protection.

Conflict of interest and financial disclosure statement

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Appendices A and B. Supplementary material

Supplementary material to this article can be found online at https://doi.org/10.1016/j.jeem.2022.102782.
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Further reading


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